

ARTICLE OPEN



Who benefits most from expectancy effects? A combined neuroimaging and antidepressant trial in depressed older adults

Sigal Zilcha-Mano¹✉, Meredith L. Wallace², Patrick J. Brown³, Joel Sneed⁴, Steven P. Roose³ and Bret R. Rutherford³

© The Author(s) 2021

Depressed patients' expectations of improvement drive placebo effects in antidepressant clinical trials, yet there is considerable heterogeneity in the magnitude of expectancy effects. The present study seeks to identify those individuals who benefit most from expectancy effects using baseline neuroimaging and cognitive measures. Older adult outpatients diagnosed with major depressive disorder (MDD) participated in a prospective, 8-week clinical trial in which expectancy was experimentally manipulated and its effects on depression outcome measured. Based on the literature, we selected a priori 12 cognitive and brain-based variables linked to depression and expectancy, together with demographic variables, and incorporated them into a combined moderator. The combined moderator was developed as a weighted combination of the individual moderators, and was used to identify individuals who benefited most from expectancy effects. The combined moderator was found to predict differential change in depression severity scores between the high- vs. low-expectancy groups with a medium-size effect (Spearman effect size: 0.28). While at the sample level no expectancy effect was found, the combined moderator divided older adults with MDD into those who did and those who did not improve as the result of expectancy manipulation, with those benefiting from the manipulation showing greater processing speed, executive function, and frontostriatal white matter tract integrity. The findings suggest that it is possible to identify a subgroup of older adult individuals with MDD for whom expectancy manipulation results in greater antidepressant treatment response, supporting a precision medicine approach. This subgroup is characterized by distinct cognitive dysfunction and neuroimaging impairments profiles.

Translational Psychiatry (2021)11:475; <https://doi.org/10.1038/s41398-021-01606-1>

INTRODUCTION

Placebo effects are among the most impactful and consistent phenomena in medicine, and they are particularly prominent in major depressive disorder (MDD), the leading cause of disability worldwide [1]. The increase in trial failures in the past decades, which makes it increasingly difficult and expensive to develop effective drugs, is attributed largely to the rising placebo response rather than to declining antidepressant medication (ADM) response [2]. It is imperative to develop methods of minimizing placebo response in antidepressant randomized control trials (RCT) to allow the valid evaluation of new ADMs [2]. At the same time, harnessing the placebo effect in clinical practice can benefit patients. Better mechanistic understanding of placebo effects, including the characteristics of individuals who most benefit from them, has the potential to facilitate achievement of these two important complementary goals.

Patient expectancy, an individuals' belief about whether and how much they will improve as a result of treatment [3], is a key mechanism underlying the placebo effect [2, 4, 5]. Meta-analyses suggest that expectancy has a large effect in ADM trials, as manifested in smaller ADM effects in placebo-controlled than in

open or active comparator trials [6–8], especially as the probability of receiving placebo vs. ADM increases [9, 10]. Our group has developed a methodology to experimentally manipulate expectancy effects prospectively, by randomizing individuals to a high-expectancy group (open trial with a 100% chance of receiving ADM) vs. a low-expectancy group (placebo-controlled trial, where the chances of receiving ADM are lower) [2]. Using this approach, we have reported that it is feasible to manipulate expectancy in young adults and that depressed individuals randomized to high expectancy conditions experience more symptomatic improvement [4]. Findings also demonstrate that in young adults, gains in expectancy during treatment result in subsequent symptom reductions [11].

Expectancy effects may be weaker or more variable in older adults with MDD, as our past studies failed to successfully manipulate expectancy based on the probability of receiving active medication in an antidepressant trial [4, 12]. Older adults with MDD are a population of great interest in identifying moderators of expectancy-based placebo effects because by virtue of brain aging they exhibit variability in cognitive (e.g., memory and executive function) and neural (e.g., integrity of

¹Department of Psychology, University of Haifa, Mount Carmel, 31905 Haifa, Israel. ²University of Pittsburgh Department of Psychiatry, Statistics, and Biostatistics, Pittsburgh, PA, USA. ³Columbia University Vagelos College of Physicians and Surgeons, New York State Psychiatric Institute, New York, NY, USA. ⁴Queens College of the City University of New York, New York, NY, USA. ✉email: sigalzil@gmail.com

Received: 30 March 2021 Revised: 27 July 2021 Accepted: 12 August 2021

Published online: 15 September 2021

frontostriatal tracts, white matter hyperintensities) markers that may be highly relevant to expectancy [13, 14]. Consistent with this possibility, we recently found that decreased processing speed interfered with expectancy effects in older adults with MDD. There was a trend in the data suggesting that other neurocognitive features may be relevant moderators of expectancy, but the effect sizes for each were small and the individual moderation effects were non-significant [12].

To address this common situation in many RCTs where a set of important moderators is identified, but each one is too small to explain the heterogeneity between individuals, a combined moderator approach has recently been proposed [15–17]. By combining multiple weak moderators into a single stronger moderator of the expectancy effect, a clinically useful index can emerge. Indeed, most studies have failed to distinguish those who do from those who do not benefit from expectancy and placebo effects [18–20]. These generally focus on single clinical and demographical characteristics: short episode duration, few previous episodes, good response to antidepressant treatment, low overall symptom severity [20], gender [19], age [21], and education [22], rather than combinations of variables or brain-based measures. A combined moderator could amplify the effects of weaker, individual moderators. Moreover, each individual moderate alone may provide conflicting treatment indications for a given individual. For example, if individuals with lower education benefit from high expectancy while individuals with high symptom severity show less benefit from expectancy effect, there is no practical guidance for an individual with both lower education and higher symptom severity. Our group has recently demonstrated the benefits of combining different moderators for the purpose of identifying older adults with MDD who may respond to placebo [23].

In the present study, we quantified cognitive and brain-based variables related to expectancy effects and depression to identify a subpopulation of older adults with MDD who benefit most from experimental manipulations of expectancy. We analyzed data from an RCT in which expectancy was experimentally manipulated by randomizing individuals to low- vs. high-expectancy conditions. Twelve potential moderators were chosen a priori based on the literature [24] focusing on cognitive and neuroimaging variables to create a combined moderator of expectancy effect. Specifically, we focused on cognitive performance deficits and white matter hyperintensities on T2-weighted magnetic resonance imaging because they are common in late-life depression, are associated with poor outcomes [13, 14, 25], and have been hypothesized to serve as the mechanisms underlying poorer expectancy effect [12]. We also included education [23] and age [4, 22] because of previous research showing that these variables are associated with placebo effect. Prior to the expectancy manipulation in this RCT, we administered a battery of

neuropsychological tests focusing on executive function, complemented by structural MRI and diffusion tensor imaging (DTI), which served as potential moderators.

METHODS

Participants

The study was conducted at the Clinic for Aging, Anxiety, and Mood Disorders at the New York State Psychiatric Institute (NYSPI). All procedures were approved by the NYSPI Institutional Review Board, and registered on Clinicaltrials.gov (NCT01931202). Eligible participants were men and women aged 60–90 years old, who met Diagnostic and Statistical Manual IV (DSM-IV) [26] criteria for non-psychotic MDD, had a 24-item Hamilton Rating Scale for Depression (HRSD) [27] score ≥ 16 , were right-handed, gave informed consent, and complied with study procedures.

Clinical trial design

Study procedures are described in a previous report [12]. Briefly, 108 patients were enrolled in an 8-week antidepressant clinical trial experimentally manipulating expectancy (Fig. 1). At baseline, patients underwent an initial evaluation to assess eligibility, had pre-randomization psychiatric symptoms and neurocognitive performance measured, and underwent MRI scanning. Next, participants' expectancy of improvement was experimentally manipulated by randomization to open administration of ADM (i.e., 100% probability of active medication, high expectancy condition) or placebo-controlled administration of ADM (i.e., perceived 50% probability of active medication, low-expectancy condition). This experimental procedure has been successful in manipulating the expectancy effect in multiple previous studies [4, 12]. Participants receiving open medication began either escitalopram or duloxetine (depending on their past treatment history), while those in the placebo-controlled condition were randomized to medication or placebo in a 6:1 ratio favoring medication. Clinicians and participants were aware of group assignment (i.e., open vs. placebo-controlled) but blinded to treatment assignment within the placebo-controlled group, whereas outcome assessors were blinded to both group and specific treatment assignment.

Measures

Neurocognitive tasks. (a) Stroop Color-Word Test, measuring response inhibition, adjusted for age and education [28], (b) Digit Symbol subtest from the WAIS-III [29], and (c) Initiation/Perseveration (I/P) subtest of the Mattis Dementia Rating Scale (DRS) [30].

Neuroimaging procedures. A GE Discovery MR750 3.0 Tesla whole-body scanner (GE Medical Systems, Waukesha, Wisconsin) and 48-channel head coil were used. A 3-plane localizer (scout) was used to determine patient position, followed by T1-weighted (FSPGR), T2 fluid-attenuated inversion recovery (FLAIR), and DTI scans. T2 FLAIR scans quantified whole-brain WMH volume in cm. DTI data were processed using FMRIB Software Library (FSL) version 6.0.1 (Oxford, UK) and analyzed with tract-based spatial statistics (TBSS). These methods provided mean fractional anisotropy (FA) and mean diffusivity (MD) values for the superior longitudinal fasciculus (SLF) and uncinate fasciculus (UNC) for each individual. Detailed descriptions of neuroimaging procedures appear in the Supplementary Materials.

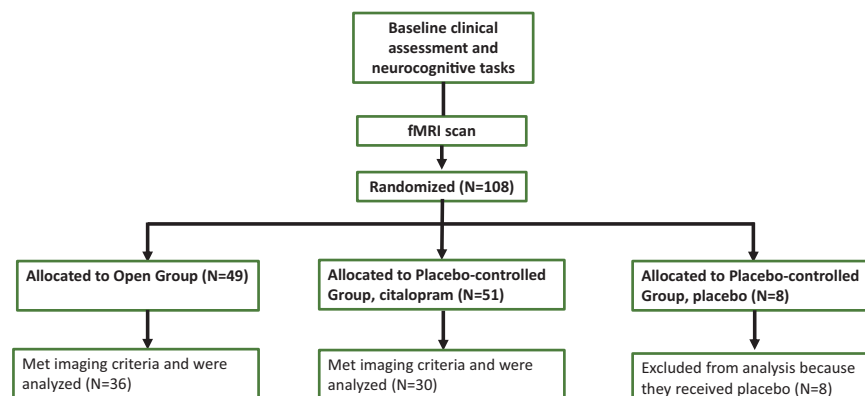


Fig. 1 Diagram flow. Diagram flow of individuals participating in the randomized controlled trial. *N* refers to the number of individuals.

Statistical analyses

We adopted a combined moderator approach to address the problem of weak individual moderator effects in clinical trials research [16, 17]. All moderators were standardized. We first created a new data set from all possible pairs of a patient in the high-expectancy and a patient in the low-expectancy conditions ($n_1 \times n_2$ pairs, where n_i is the number of patients in condition i , $i=1,2$). For each pair, we calculated the average of their moderator values (for each moderator), and the difference in their estimated HRSD slopes. The slope of each patient was estimated using a mixed-effect regression model, with random slope and intercept for each patient, based on a linear trend in log of week. For each potential moderator, we computed the non-parametric Spearman correlations in the new data between the differences in HRSD slopes and the covariate average across all pairs. Non-parametric Spearman correlation was used to allow for non-normal moderations and to reduce the potential influence of outliers in the data.

Second, we created the combined moderator, which is an optimally weighted combination of individual moderators. The weight assigned to each moderator was estimated using a LASSO regression with the glmnet package in R. The dependent variable in the Lasso regression was the slope difference of each pair, and the predictors were the averages of the potential moderators. The Lasso regression is a linear regression, but it uses a penalty on coefficient absolute values, which shrinks their estimate. The purpose of the shrinkage is to avoid overfitting. Unimportant predictors are expected to shrink to zero. The estimated coefficients (standardized to sum to 1 in absolute values) are used as weights for calculating the combined moderator score of each individual. In this way, the combined moderator represents an optimally weighted linear combination of the individual moderator scores. The resulting effect is measured by the correlation of the averages of the combined moderator (M) and the paired slope differences (in the paired data), denoted by $Cor(M)$. We used bootstrapping to obtain a confidence interval. We resampled with replacement patients, separately from each condition, and applied all the procedures above, resulting in 1000 estimates of the correlation. The 2.5 and 97.5 quantiles of the estimates served as confidence intervals. Finally, we assessed the discriminative utility of the combined moderator, using a linear regression of expectancy condition, the combined moderator, and the combined moderator by condition interaction on treatment outcome.

RESULTS

Demographic and clinical characteristics

Of the 108 patients participating in the RCT, 8 received placebo and therefore were excluded from the analyses. A total of 66 individuals who received ADM underwent baseline MRI scanning, creating the effective sample for this analysis. Of these patients, 36 were randomized to the high expectancy and 30 to the low-expectancy group (Fig. 1). No significant differences in demographic data, baseline clinical characteristics or outcome were found between participants who were and were not scanned (Table S1 in the online supplements) or between participants in the Placebo-controlled and Open groups (Table S2 in the online supplements). The two groups did not differ significantly also in pre-treatment depression scores (23.7 vs. 22.5 in the Placebo-controlled vs. Open group, respectively; $p = .38$). The two groups did not differ significantly in their HRSD slopes from pre-treatment to post treatment (-0.27 vs. 0.76 in the Placebo-controlled vs. Open group, respectively; $p = 0.096$), suggesting the need to identify potential moderators to explain the variability in this effect.

Creating the combined moderator

Across the entire dataset, there were only 15 missing observations (6 in WMH and 9 in education years) in the 12-variables of interest. The few missing observations were imputed using MissForest. As expected, the effect for the individual moderators was small and included both positive and negative correlations (Table 1). Positive correlations mean that higher levels of the moderator were associated with less improvement in treatment outcome in the high vs. low-expectancy condition. The largest individual moderators were fractional anisotropy (FA) values in the anterior thalamic radiations (left and right), followed by left superior longitudinal

Table 1. Individual moderator effect sizes and their weights in the combined moderator.

	Correlation	Weight
WAIS digit symbol	-0.01	-0.09
Stroop	0.14	0.16
WMH	-0.05	0.00
Mattis DRS initiation/perseveration	-0.06	-0.01
DTI: Anterior thalamic radiation L	0.16	0.03
DTI: Anterior thalamic radiation R	0.16	0.13
DTI: Superior longitudinal fasciculus L	0.14	0.19
DTI: Superior longitudinal fasciculus R	0.07	-0.10
DTI: Uncinate fasciculus L	0.01	-0.15
DTI: Uncinate fasciculus R	-0.08	-0.05
Age	-0.03	-0.04
Education	-0.02	0.06
Combined moderator	0.28	

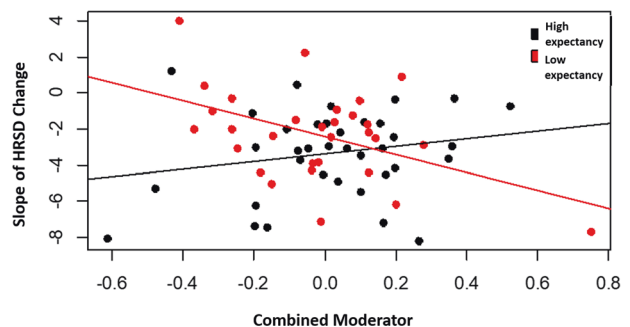


Fig. 2 Combined moderator by expectancy interaction. The linear interaction of the combined moderator and expectancy condition in predicting the slope of symptom reduction.

fasciculus FA and Stroop Color-Word score, all showing a positive correlation. The largest negative correlations were found for right uncinate fasciculus FA and Mattis DRS Initiation/Perseveration subscale score.

The combined moderator effect

The combined moderator had a larger effect size than any individual moderator (effect size = 0.28 (95% CI: .27, .67 vs. $.01 \leq R \leq 0.16$ for the individual moderators). The linear interaction of the combined moderator and expectancy condition is shown in Fig. 2. When the combined moderator is lower than the cross-point (0.14), the high-expectancy condition showed more symptom reduction than the low-expectancy condition (Cohen's $d = 0.58$; -7.41 vs. -4.46 points reduction in HRSD for the high- vs. low-expectancy conditions). When the combined moderator is higher than the cross-point, the differences were not substantial (Cohen's $d = 0.12$; -7.17 vs. -8.03 points reduction in HRSD for the high- vs. low-expectancy conditions). The effect sizes after removing each moderator separately appear in Table S3 of the online supplements. As a post hoc analysis, we repeated this procedure adding sex, ethnicity, and race. Findings appear in Table S4 of the online supplements.

DISCUSSION

The present findings reveal that some depressed older adults benefited more than others from the experimental manipulation of expectancy during antidepressant treatment. Specifically,

participants with a lower combined moderator score experienced significantly greater reductions in HRSD scores in the high- vs. low-expectancy conditions compared to other participants who benefited substantially less from the expectancy manipulation. These two distinct subgroups were concealed when focusing on each potential moderator separately: only small effects ranging in size between $0.01 \leq R \leq 0.16$ were observed for each separate neurocognitive moderator. The combined effect was much larger ($R = 0.28$), and was able to differentiate between those who benefit from expectancy manipulation (with an effect size of Cohen's $d = 0.58$) and those who do not (Cohen's $d = 0.12$).

Contingent upon future external validation and prospective research, the combined moderator identified in the present study can be used as a clinically useful index to predict who is likely to benefit from an intervention to enhance expectancy. That is, upon assessing a new patient and having access to the measurements contained in the combined moderator, these values could be used to obtain an estimate of the patient's likelihood of benefiting from a high expectancy compared to a lower expectancy intervention. This may have implications for future randomized trial inclusion criteria as well as for daily clinical practice. The index may be used to direct efforts for developing new drugs for the populations that are less expected to show an expectancy effect, which may facilitate drug/placebo signal detection by minimizing placebo response. That is, this type of analysis may be useful in predicting a randomized controlled trial participant's likelihood of demonstrating a high placebo response. Such information may be useful in designing selection criteria and/or stratifying samples so as to maximize signal detection for novel therapeutic agents. At the same time, the use of the combined moderator may allow targeting of patients who are likely to benefit from expectancy effects with enhanced psychoeducation about the potential benefits of treatment. Based on the index, the attending physician may be able to use expectancy augmentations techniques in combination with the ADM, for example, by further informing the patient about the expected positive effects of a given drug.

The present findings can also shed light on the critical capabilities required for showing an expectancy effect, answering questions that have been of great interest in the empirical literature on placebo effect. The findings may suggest that of the capabilities evaluated here, executive functioning and frontostriatal tract integrity, are especially critical. Intact executive functioning may be critical in reappraising the responses to an event according to new information presented in the world (in this case, verbal information regarding the probability of receiving an active drug). Such new information is then processed and evaluated through circuits implicated in generating an expectancy effect. Reduced integrity of the frontostriatal tract may interrupt the modulation of limbic and striatal structures necessary to reduce depressive symptoms as a result of the expectancy manipulation. This potential mechanistic explanation for the expectancy effect is consistent with previous findings demonstrating the importance of executive functioning in producing expectancy effects and in symptom reduction in older adults with MDD [25]. It is also consistent with previous findings suggesting that DTI may be implicated in ADM non-response in late life depression [12].

The approach used in the present study goes beyond previous research, by combining distinct moderators to better capture the richness and complexity of the neurocognitive capacities that are needed to benefit from the expectancy effect. It has the potential to leverage current research on placebo responders by combining the weak moderators identified so far in a way that captures their specific importance. The present findings are consistent with previous ones demonstrating the promise of this approach in translating heterogeneity in clinical outcomes into personalized recommendations [31–33]. To be validated, the algorithm must be prospectively tested. One potential test is by randomizing

individuals to either (a) receive randomly high- vs. low-expectancy manipulation, or (b) be assigned to high- vs. low-expectancy manipulation using the algorithm specified in the current study (namely, stratifying individuals by whether they are above or below the cut-point of the combined moderator). We expect the assignment by algorithm to vastly outperform the random assignment. Future research should further test how weights could be adjusted to tailor the algorithm to diverse populations.

The most important limitation of the current study is that our sample size was restricted by the unique characteristics of the sample that underwent the experimental manipulation of expectancy and required each patient to have rich and detailed cognitive and neuroimaging characteristics. Future studies should use a larger sample and test the validity of the prediction on an external sample. Until then, the findings should be regarded as exploratory. Notwithstanding these limitations, the findings identify a subpopulation of older adults with MDD who benefited from expectancy manipulation: those individuals with intact executive functioning (enabling reappraising responses based on the new expectancy-related information arriving), as well as less reduced integrity of the frontostriatal tract (enabling the modulation of limbic and striatal structures). These findings have the potential to greatly advance our understanding of the pathogenesis of late-life depression, and shed light on the biology of the expectancy effect in antidepressant response. The findings further hold the potential for improving the efficacy of treatment of late life depression through more precise treatment selection, focusing on psychoeducational interventions [34] and on interventions aimed at improving response inhibition [35].

REFERENCES

- Friedrich MJ. Depression is the leading cause of disability around the world. *JAMA*. 2017;317:1517.
- Rutherford BR, Roose SP. A model of placebo response in antidepressant clinical trials. *Am J Psychiatry*. 2013;170:723–33.
- Devilley GJ, Borkovec TD. Psychometric properties of the credibility/expectancy questionnaire. *J Behav Ther Exp psychiatry*. 2000;31:73–86.
- Rutherford BR, Wall MM, Brown PJ, Choo TH, Wager TD, Peterson BS, et al. Patient expectancy as a mediator of placebo effects in antidepressant clinical trials. *Am J Psychiatry*. 2017;174:135–42.
- Atlas LY, Wager TD. How expectations shape pain. *Neurosci Lett*. 2012;520:140–8.
- Greenberg RP, Bornstein RF, Greenberg MD, Fisher S. A meta-analysis of antidepressant outcome under "blinder" conditions. *J Consulting Clin Psychol*. 1992;60:664–9.
- Rutherford BR, Sneed JR, Roose SP. Does study design influence outcome? *Psychother Psychosom*. 2009;78:172–81.
- Sneed JR, Rutherford BR, Rindskopf D, Lane DT, Sackeim HA, Roose SP. Design makes a difference: a meta-analysis of antidepressant response rates in placebo-controlled versus comparator trials in late-life depression. *Am J Geriatr Psychiatry*. 2008;16:65–73.
- Khan A, Kolts RL, Thase ME, Krishnan KRR, Brown W. Research design features and patient characteristics associated with the outcome of antidepressant clinical trials. *Am J Psychiatry*. 2004;161:2045–9.
- Sinyor M, Levitt AJ, Cheung AH, Schaffer A, Kiss A, Dowlati Y, et al. Does inclusion of a placebo arm influence response to active antidepressant treatment in randomized controlled trials? Results Pool meta-analyses. 2010;71:270–9.
- Zilcha-Mano S, Brown PJ, Roose SP, Cappetta K, Rutherford BR. Optimizing patient expectancy in the pharmacologic treatment of major depressive disorder. *Psychological Med*. 2019;49:2414–20.
- Rutherford BR, Choi J, Slifstein M, O'Boyle K, Abi-Dargham A, Brown PJ, et al. Slowed processing speed disrupts patient expectancy in late life depression. *Am J Geriatr Psychiatry*. 2020;265:439–444.
- Butters MA, Whyte EM, Nebes RD, Begley AE, Dew MA, Mulsant BH, et al. The nature and determinants of neuropsychological functioning in late-life depression. *Arch Gen psychiatry*. 2004;61:587–95.
- Lavretsky H, Lesser IM, Wohl M, Miller BL, Mehlinger CM. Clinical and neurobiologic features associated with chronicity in late-life depression. *Am J Geriatr Psychiatry*. 1999;7:309–16.

15. Wallace ML, Smagula SH. The promise and challenges of using combined moderators methods to personalize mental health treatment. *Am J Geriatr Psychiatry*. 2018;26:678–679.
16. Kraemer HC. Discovering, comparing, and combining moderators of treatment on outcome after randomized clinical trials: a parametric approach. *Stat Med*. 2013;32:1964–73.
17. Wallace ML, Frank E, Kraemer HC. A novel approach for developing and interpreting treatment moderator profiles in randomized clinical trials. *JAMA psychiatry*. 2013;70:1241–7.
18. Fava M, Evins AE, Dorer DJ, Schoenfeld DA. The problem of the placebo response in clinical trials for psychiatric disorders: culprits, possible remedies, and a novel study design approach. *Psychother Psychosom*. 2003;72:115–27.
19. Mallinckrodt CH, Zhang L, Prucka WR, Millen BA. Signal detection and placebo response in schizophrenia: parallels with depression. *Psychopharmacol Bull*. 2010;43:53–72.
20. Brown WA, Johnson MF, Chen M-G. Clinical features of depressed patients who do and do not improve with placebo. *Psychiatry Res*. 1992;41:203–14.
21. Papakostas GI, Fava M. Does the probability of receiving placebo influence clinical trial outcome? A meta-regression of double-blind, randomized clinical trials in MDD. *Eur Neuropsychopharmacol*. 2009;19:34–40.
22. Buitelaar JK, Sobanski E, Stieglitz RD, Dejonckheere J, Waechter S, Schäuble B. Predictors of placebo response in adults with attention-deficit/hyperactivity disorder: data from 2 randomized trials of osmotic-release oral system methylphenidate. *J Clin psychiatry*. 2012;73:1097–102.
23. Zilcha-Mano S, Roose SP, Brown PJ, Rutherford BR. A machine learning approach to identifying placebo responders in late-life depression trials. *Am. J. Geriatr Psychiatry*. 2018;26:669–677.
24. Rutherford BR, Taylor WD, Brown PJ, Sneed JR, Roose SP. Biological aging and the future of geriatric psychiatry. *J Gerontol Ser A: Biomed Sci Med Sci*. 2017;72:343–52.
25. Tunvirachaisakul C, Gould RL, Coulson MC, Ward EV, Reynolds G, Gathercole RL, et al. Predictors of treatment outcome in depression in later life: a systematic review and meta-analysis. *J Affect Disord*. 2018;227:164–82.
26. American Psychiatric Association. *Diagnostic and Statistical Manual of Mental Disorders—Text Revision*. 4. Washington DC: American Psychiatric Press; 2000.
27. Hamilton MAX. Development of a rating scale for primary depressive illness. *Br J Soc Clin Psychol*. 1967;6:278–96.
28. Golden CJ, Freshwater SM. Stroop color and word test kit for adults, Wood Dale. Stoelting; 2002.
29. Wechsler D. *WAIS-iii*. San Antonio, TX: Psychological corporation; 1997.
30. Mattis S. *Dementia rating scale: professional manual*. Psychological Assessment Resources, Odessa. Incorporated; 1988.
31. Wallace ML, Banihashemi L, O'Donnell C, Nimgaonkar VL, Kodavali C, McNamee R, et al. Using optimal combined moderators to define heterogeneity in neural responses to randomized conditions: Application to the effect of sleep loss on fear learning. *Neuroimage*. 2018;181:718–27.
32. Smagula SF, Wallace ML, Anderson SJ, Karp JF, Lenze EJ, Mulsant BH, et al. Combining moderators to identify clinical profiles of patients who will, and will not, benefit from aripiprazole augmentation for treatment resistant late-life major depressive disorder. *J Psychiatr Res*. 2016;81:112–8.
33. Smagula SF, Freedland KE, Steinmeyer BC, Wallace MJ, Carney RM, Rich MW. Moderators of response to cognitive behavior therapy for major depression in patients with heart failure. *Psychosom Med*. 2019;81:506–12.
34. Constantino MJ, Ametrano RM, Greenberg RP. Clinician interventions and participant characteristics that foster adaptive patient expectations for psychotherapy and psychotherapeutic change. *Psychotherapy*. 2012;49:557–69.
35. Zhao X, Chen L, Maes JHR. Training and transfer effects of response inhibition training in children and adults. *Developmental Sci*. 2018;21:e12511.

ACKNOWLEDGEMENTS

We are thankful for the helpful feedback by Dr. Helena Kraemer on the statistical analyses. This study was funded by NIMH R01 MH102293 (to Dr. Rutherford). Work on this paper was supported by The U.S.-Israel Binational Science Foundation (BSF) grant 2017263 (to Dr. Zilcha-Mano).

AUTHOR CONTRIBUTIONS

The conceptualization and design were done by SZM and BRR. Data analyses were conducted by SZM in consultation with MLW. Interpretations of the findings were done by SZM, MLW, PJB, SPR, and BRR. SZM wrote the first draft of the manuscript. MLW, PJB, SPR, and BRR reviewed the manuscript and provided feedback.

COMPETING INTERESTS

Drs. Zilcha-Mano and Rutherford had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. Drs. Zilcha-Mano, Brown, Sneed, Rutherford, and Roose have no disclosures or conflicts of interest to report. Dr. Wallace receives statistical consulting fees from Noctem and grant funding from the National Institute on Aging, both unrelated to this work. This paper has not been previously presented.

ADDITIONAL INFORMATION

Supplementary information The online version contains supplementary material available at <https://doi.org/10.1038/s41398-021-01606-1>.

Correspondence and requests for materials should be addressed to Sigal Zilcha-Mano.

Reprints and permission information is available at <http://www.nature.com/reprints>

Publisher's note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.



Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons license, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons license and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this license, visit <http://creativecommons.org/licenses/by/4.0/>.

© The Author(s) 2021